Bayesian Double Machine Learning for Causal Inference

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My Research Interests

Econometrics



Causal Inference, Spillovers, Measurement Error, Model Selection, Bayesian Inference

Applied Work

Childhood Lead Exposure, Pawn Lending in Mexico City, Colombian Civil Conflict

Overview of Today's Talk

- Causal inference is hard, especially when there are many controls.
- Bayesian approach is appealing, but doesn't work out-of-the-box
- Find a way to combine the advantages of Bayes with good Frequentist properties (bias / variance / coverage probability)
- Related to Frequentist literature on "Double Machine Learning" but aims to improve on finite-sample performance.
- ▶ Workshop on Bayesian Causal Inference this Friday: email me for a link!

The Problem / Model

$$Y_i = \alpha D_i + X'_i \beta + \varepsilon_i, \quad \mathbb{E}[\varepsilon_i | D_i, X_i] = 0, \quad i = 1, \dots, n$$

- Learn effect α of treatment D_i (not necessarily binary)
- Selection-on-observables: p-vector of controls X_i
- ▶ OLS: unbiased and consistent estimator of α , but noisy if p is large
- ▶ Drop control $X_i^{(j)}$ that is correlated with $D_i \Rightarrow$ biased estimate of α if $\beta^{(j)} \neq 0$.

Naïve Shrinkage Estimator: Ridge Regression

Assume everything de-meaned, X scale-normalized

Unique, closed-form solution even if p > n

$$\begin{bmatrix} \widehat{\alpha}_{\mathsf{naive}} \\ \widehat{\beta}_{\mathsf{naive}} \end{bmatrix} = \begin{bmatrix} \begin{pmatrix} D'D & D'X \\ X'D & X'X \end{pmatrix} + \begin{pmatrix} 0 & 0'_p \\ 0_p & \lambda \mathbb{I}_p \end{pmatrix} \end{bmatrix}^{-1} \begin{pmatrix} D'Y \\ X'Y \end{pmatrix}, \quad \lambda \equiv \frac{\sigma_{\varepsilon}^2}{\sigma_{\beta}^2}.$$

Frequentist Interpretation

Minimize
$$(Y - \alpha D - X\beta)'(Y - \alpha D - X\beta) + \lambda\beta'\beta$$

Bayesian Interpretation

Posterior mean: σ_{ε} known, flat prior on α , independent Normal $(0, \sigma_{\beta}^2)$ priors on β_j

Regularization-Induced Confounding (RIC)

Term coined by Hahn et al. (2018)

If $\lambda > 0$, bias from correlation between D and residuals:

$$\begin{split} \mathsf{Bias}(\widehat{\alpha}_{\mathsf{naive}}) &= \widehat{\omega}' \left[\mathbb{I}_p - (R + \lambda \mathbb{I}_p)^{-1} R \right] \beta \\ \mathsf{Var}(\widehat{\alpha}_{\mathsf{naive}}) &= \sigma_{\varepsilon}^2 \left[(D'D)^{-1} + \widehat{\omega}' (R + \lambda \mathbb{I}_p)^{-1} R (R + \lambda \mathbb{I}_p)^{-1} \widehat{\omega} \right] \end{split}$$

Notation

$$\widehat{\omega}_j = (D'D)^{-1}D'X_j, \quad \widehat{E}_j = X_j - \widehat{\omega}_j X_j, \quad R = \widehat{E}'\widehat{E}$$

Problem

For $\lambda > 0$, bias depends crucially on $\hat{\omega}$ and β ; strong confounding \Rightarrow large bias

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Adding a First-Stage

Just a Projection

$$Y = \alpha D + X'\beta + \varepsilon, \quad \mathbb{E}[\varepsilon|X, D] = 0$$
$$D = X'\gamma + V, \quad \mathbb{E}[V|X] = 0$$

Implied by Casual Assumption

$$\mathsf{Cov}(arepsilon, V) = \mathsf{Cov}(arepsilon, D - \mathsf{X}'\gamma) = \mathsf{Cov}(arepsilon, D) - \mathsf{Cov}(arepsilon, \mathsf{X}')\gamma = 0.$$

Idea

Maybe adding this regression allows us to learn the degree of counfounding.

Adding the D on X regression has no effect!

"Bayes Ignorability" - Linero (2023; JASA)

Bayes' Theorem $\pi(\theta|Y, D, X) \propto f(Y, D|X, \theta) \times \pi(\theta)$ $Cov(\varepsilon, V) = 0 \Rightarrow no common parameters!$ $f(Y, D|X, \theta) = f(Y|D, X, \theta)f(D|X, \theta) = f(Y|D, X, \alpha, \beta, \sigma_{\varepsilon}^{2}) \times f(D|X, \gamma, \sigma_{V}^{2})$

Problem

Unless prior treats β and γ as dependent, adding the D on X regression has no effect!

Our Solution: Bayesian Double Machine Learning (BDML)

From Structural to Reduced Form

$$Y_i = \alpha D_i + X'_i \beta + \varepsilon_i = X'_i (\alpha \gamma + \beta) + (\varepsilon_i + \alpha V_i) = X'_i \delta + U_i$$

$$\begin{array}{ll} Y_i = X'_i \delta + U_i & \begin{bmatrix} U_i \\ V_i \end{bmatrix} \\ X_i \sim \operatorname{Normal}_2(0, \Sigma), \quad \Sigma = \begin{bmatrix} \sigma_{\varepsilon}^2 + \alpha^2 \sigma_V^2 & \alpha \sigma_V^2 \\ \alpha \sigma_V^2 & \sigma_V^2 \end{bmatrix}$$

BDML Algorithm

- 1. Place "standard" priors on reduced form parameters (δ, γ, Σ)
- 2. Draw from posterior $(\delta, \gamma, \Sigma)|(X, D, Y)|$
- 3. Posterior draws for $\Sigma \implies$ posterior draws for $\alpha = \sigma_{UV}/\sigma_V^2$

BDML versus Frequentist Double Machine Learning (FDML)

e.g. Chernozhukov et al. (2018; Econometrics J.)

FDML Optimizes

Plug in "Machine Learning" estimators of reduced form parameters: $(\hat{\delta}_{ML}, \hat{\gamma}_{ML})$

$$\widehat{\alpha}_{\mathsf{FDML}} = \frac{\sum_{i=1}^{n} (Y_i - X'_i \widehat{\delta}_{\mathsf{ML}}) (D_i - X'_i \widehat{\gamma}_{\mathsf{ML}})}{\sum_{i=1}^{n} (D_i - X'_i \widehat{\gamma}_{\mathsf{ML}})^2}$$

BDML Marginalizes

Posterior for α averages over posterior uncertainty about γ and δ

Theoretical Results

$$egin{aligned} Y_i &= X_i'\delta + U_i & \left[egin{aligned} U_i \ V_i \end{aligned}
ight] ert X_i &\sim ext{Normal}_2(0, \Sigma) \ D_i &= X_i'\gamma + V_i & \left[egin{aligned} V_i \ V_i \end{aligned}
ight] ert X_i &\sim ext{Normal}_2(0, \Sigma) \end{aligned}$$

 $egin{aligned} &\pi(\Sigma,\delta,\gamma) \propto \pi(\Sigma)\pi(\delta)\pi(\gamma) \ &\Sigma \sim ext{Inverse-Wishart}(
u_0,\Sigma_0) \ &\delta \sim ext{Normal}_p(0,\mathbb{I}_p/ au_\delta) \ &\gamma \sim ext{Normal}_p(0,\mathbb{I}_p/ au_\gamma) \end{aligned}$

Naïve Approach

Analogous but with single structural equation and $\beta \sim \text{Normal}(0, \mathbb{I}_p/\tau_\beta)$

Asymptotic Framework

Fixed true parameters ($\Sigma^*, \delta^*, \gamma^*$); $n \to \infty$ (large sample); $p \to \infty$ (many controls)

Our asymptotic framework ensures bounded R-squared.

Rate Restrictions

(i) sample size dominates # of controls: p/n
ightarrow 0

(ii) sample size dominate prior precisions: $\tau/n \to 0$

(iii) precisions of same order as # controls: $\tau \asymp p$

Regularity Conditions

(i)
$$p < n$$

(ii) $\operatorname{Var}(X) \equiv \Sigma_X$ "well-behaved" as $p \to \infty$
(iii) $\lim_{p\to\infty} \sum_{j=1}^{p} (\delta_j^*)^2 < \infty$, $\lim_{p\to\infty} \sum_{j=1}^{p} (\gamma_j^*)^2 < \infty$
(iv) iid errors/controls, $\mathbb{E}(X_i) = 0$, finite & p.d. Σ^*



Selection Bias in the Limit

When p and n are large, what are our implied beliefs about selection bias?

$$\mathsf{SB} \equiv \left[\mathbb{E}(Y_i | D_i = 1) - \mathbb{E}(Y_i | D_i = 0)\right] - \alpha = \left[\mathbb{E}(X_i | D_i = 1) - \mathbb{E}(X_i | D_i = 0)\right]' \beta$$

Naïve Model

Degenerate prior centered at zero:
$$SB = \frac{\gamma' \Sigma_X \beta}{\sigma_V^2 + \gamma' \Sigma_X \gamma} \rightarrow_{\rho} 0$$

BDML

Non-degenerate prior centered at zero:

$$\mathsf{SB} o_{p} rac{\sigma_{UV}}{\sigma_{V}^{2} + \gamma' \Sigma_{X} \gamma}$$

. .

Summary of Asymptotic Results

Consistency

Naïve, BDML and FDML all provide consistent estimators of α .

Asymptotic Bias

BDML and FDML have bias of order p^2/n^2 compared to p/n for Naïve.

\sqrt{n} -Consistency

Naïve requires $p/\sqrt{n} \rightarrow 0$; BDML and FDML require only $p/n^{3/4} \rightarrow 0$.

Why do we focus on variance?

Bias dominates: if $p/\sqrt{n} \rightarrow 0$, all three have the same AVAR.

Simulation Experiment

$$\begin{cases} X_i \}_{i=1}^n \sim \text{iid Normal}_p(0, \mathbb{I}_p) \\ Y_i = \alpha D_i + X'_i \beta + \varepsilon_i \\ D_i = X'_i \gamma + V_i \end{cases} \begin{cases} (\varepsilon_i, V_i) \}_{i=1}^n \mid X \sim \text{iid Normal}_2\left(0, \text{diag}\left\{\sigma_{\varepsilon}^2, 1\right\}\right) \\ \beta \mid (X, \varepsilon, V) \sim \text{Normal}_p\left(\mu_{\beta}, \sigma_{\beta}^2 \mathbb{I}\right). \end{cases}$$

Linero's (2023) "Fixed" Design

$$\alpha = 2, \quad \gamma = \iota_p / \sqrt{p}, \quad \mu_\beta = -\gamma/2, \quad \sigma_\beta^2 = 1/p, \quad n = 200, \quad p = 100$$

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Two Versions of BDML

Both Versions

LKJ(4) Prior on Corr(U, V); Independent Cauchy(0, 2.5) priors on SD(U) and SD(V)

Basic Version

Independent Normal(0,5²) priors on the elements of δ and γ .

Hierarchical Version

- ▶ Independent Normal $(0, \sigma_{\delta}^2)$ priors on the elements of δ
- Independent Normal(0, σ_{γ}^2) priors on the elements of γ
- Independent Inverse-Gamma(2,2) priors on $\sigma_{\delta}, \sigma_{\gamma}$.

Two-Step "Plug-in" Bayesian Approaches

Preliminary Regression

 $\widehat{D}_i \equiv X'_i \widehat{\gamma}_{\text{prelim}} \leftarrow \text{estimate from Bayesian regression of } D \text{ on } X.$

HCPH (Hahn et al, 2018; Bayesian Analysis)

- 1. Bayesian linear regression of Y on $(D \widehat{D})$ and X
- 2. Estimation / inference for α from posterior for $(D \hat{D})$ coefficient.

Linero (2023; JASA)

- 1. Bayesian linear regression of Y on (D, \hat{D}, X) .
- 2. Estimation / inference for α from posterior the D coefficient.

Simulation Results – 3000 Replications



Only BDML and Linero have correct coverage (Left); Also lower RMSE (Right)

Zooming In: BDML versus Linero



Coverage of Linero & BDML-Hier comparable; BDML-Hier: shortest intervals & lowest RMSE

Thanks for listening!

Summary

- Simple, fully-Bayesian causal inference in a workhorse linear model with many controls.
- Avoids RIC; Excellent Frequentist Properties

In Progress

- More Simulations; Empirical Examples
- ► Good "default" prior choices?
- Extensions: partially linear model; treatment interactions; instrumental variables?

